1 Title

- 2 The neural and physiological substrates of real-world attention change across development.
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14 Abstract

- The ability to allocate and maintain visual attention enables us to adaptively regulate perception 15 and action, guiding strategic behaviour within complex, dynamic environments. This capacity to 16 regulate attention develops rapidly over the early years of life, and underpins all subsequent 17 cognitive development and learning. From screen-based experiments we know something about 18 how attention control is instantiated in the developing brain, but we currently understand little about 19 the development of the capacity for attention control within complex, dynamic, real-world settings. 20 To address this, we recorded brain activity, autonomic arousal and spontaneous attention patterns 21 in N=58 5- and 10-month-old infants during free play. We used time series analyses to examine 22 whether changes in autonomic arousal and brain activity anticipate attention changes or follow on 23 24 from them. Early in infancy, slow-varying fluctuations in autonomic arousal forward-predicted 25 attentional behaviours, but cortical activity did not. By later infancy, fluctuations in fronto-central 26 theta power associated with changes in infants' attentiveness and predicted the length of infants' 27 attention durations. But crucially, changes in cortical power followed, rather than preceded, infants' attention shifts, suggesting that processes after an attention shift determine how long that episode 28 will last. We also found that changes in fronto-central theta power modulated changes in arousal at 29 30 10 but not 5 months. Collectively, our results suggest that the modulation of real-world attention 31 involves both arousal-based and cortical processes but point to an important developmental transition. As development progresses, attention control systems become dynamically integrated 32 33 and cortical processes gain greater control over modulating both arousal and attention in naturalistic 34 real-world settings.
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36 Teaser

37 Understanding how our brains develop the capacity to pay attention to objects and people around 38 them is central for understanding all aspects of subsequent cognitive development and learning. However, we understand little about the mechanisms through which attention control develops 39 during early life in complex, real-world settings. Here, we examined the associations between 40 naturalistic attention patterns, arousal and neural activity, in 5- and 10-month-olds during free play 41 with toys. Early in infancy, lower-order arousal-based endogenous factors forward-predicted 42 attention; later in infancy, higher-order neural responses after the onset of a new attention episode 43 forward-predicted attention and modulated changes in arousal. Overall, our results suggest that the 44 neural and physiological substrates of real-world attention change between early and later infancy, 45 resulting in a more inter-linked system where associations between attention behaviour, cortical 46 activity, and autonomic arousal are stronger as infants develop. 47

48 **1. Introduction**

49 The ability to allocate and maintain visual attention enables the flexible regulation of perception and action

50 that is characteristic of strategic behaviour (1,2). The capacity to pay attention develops rapidly over the

51 early years of life (3), and individual differences in early attention predict long-term cognitive and clinical

52 outcomes (4,5). Recent new methodological advances such as naturalistic neuroimaging are allowing us to

53 build on previous research using lab-based behavioural experiments and animal studies.

54 The development of attention is traditionally conceptualised as the product of interactions among different 55 systems at different levels of maturity (1,6,7). Traditionally, the earliest subcomponent of attention to develop is thought to be the arousal/ alertness subcomponent, mediated via brainstem reticular activating 56 57 systems centred on the locus coeruleus (LC) and instantiated primarily via norepinephrine neurotransmitter 58 systems (6). In young infants, alertness is more readily initiated by exogenous events (8); over time, infants gain the ability to both attain and maintain an alert state even in absence of external stimulation. Areas 59 60 around the brainstem (including the LC) are thought to be some of the earliest to become functionally mature (9,10). Consequently, the relative influence of this subcomponent of attention is thought to be 61 62 strongest during early development (6).

Behaviourally, the arousal/ alertness subcomponent of attention is thought to reflect a state of anticipatory readiness, or alertness for stimulus input (6). Arousal is generally measured indirectly, via proxy measures of autonomic nervous system activity such as heart rate (11). Heart rate has been extensively studied in the context of infant attention (3,12,13). During anticipatory readiness, we know that reorientations of visual attention take place periodically, clustered around a preferred modal reorientation rate (14–18). This may reflect rhythmic activity in the central nervous system (19).

With time, it is thought that looking behaviours become increasingly modulated by higher-level executive processes that reflect the infant's internal states, motivation, comprehension, and goals (2,20,21). Behaviourally, this increase in endogenous or internally directed attention has been shown as: a developmental increase in the degree to which attentional engagement is accompanied by decreases in distractibility (22,23); an increase in selective attention as measured indirectly, using the blink reflex (24); and differences in the trajectory of how attention durations to simple vs complex stimuli change over developmental time (25).

Other research that used experimenter-controlled, screen-based tasks to examine neural correlates of attention has examined changes in the power spectral density (PSD) of EEG oscillations, in particular infants' theta (3-6Hz) rhythm, which increases during active, anticipatory, and exploratory behaviour (26– 79 31). Together, these studies suggest that the expression of theta during attention-eliciting episodes could 80 signify the engagement of neural networks related to executive attention (28,29,31). Similarly, other studies 81 have reported decreases in alpha band activity under conditions of increased attention (32,33). Both theta 82 and alpha effects are now widely known in the literature as "theta synchronization" and "alpha 83 desynchronization" (29).

How children allocate their attention in experimenter-controlled, screen-based lab tasks differs, however, from actual real-world attention in several ways (34–36). For example, the real-world is interactive and manipulable, and so how we interact with the world determines what information we, in turn, receive from it: experiences generate behaviours (37). While lab-based studies can be made interactive (e.g., 38,39), how infants actively and freely initiate and self-structure their attention remains unexplored (40).

89 The present study aims to examine developmental changes in the relationship between autonomic arousal, cortical activity, and attention in real-world settings. To do this, we first explored how naturalistic attention 90 91 patterns (measured via looking durations to play objects) from a solo play interaction change between 5 and 10 months. Then, we explored temporal relations between changes in infant arousal (measured via heart 92 93 rate) and attention episodes in typical 5- and 10-month-olds infants. Finally, we investigated changes in 94 EEG theta power relative to attention episodes, and changes in EEG theta relative to arousal (see Figure 1). 95 As attentional systems mature and brain regions become increasingly specialised (41,42), we expected to 96 see both a developmental increase in attention towards play objects and a developmental shift in the way 97 different mechanisms (i.e., arousal/ alertness vs. executive attention subsystems) drive attention. To 98 measure attention we used looking time, an approach that is known to have several limitations (12,43,44). 99 For example, we cannot differentiate based on looking time alone whether overt and covert attention are 100 coupled or decoupled (see 3.1).

101 Our first set of analyses examined attentional inertia (the phenomenon that, as individuals become 102 progressively more engaged with an object, their attention progressively increases) as a measure of internally driven attentional engagement (3,45,46). We tested whether attentional inertia influenced 103 104 attentional behaviours more strongly at 10 months compared to 5 months. To do so, we calculated both the 105 Autocorrelation Function (ACF) and the survival probability of spontaneously occurring attention episodes during play (analysis 1). The ACF allowed us to quantify the rate of change of spontaneous attention 106 107 durations. A faster rate of change would indicate lower attentional inertia. The survival probability, on the other hand, allowed us to quantify the probability between looking (i.e., paying attention) and looking away. 108 A slower decrease in the probability of an attention episode surviving would indicate increased attention 109 110 engagement and decreased distractibility by other stimuli. We hypothesised that, as slow-varying

fluctuations in endogenous interest or engagement start to influence looking behaviour more strongly over time, 10-month-old infants would show increased endogenous attention control indexed by a slower rate of change of attentiveness and slower decreases in the survival probability. We also predicted that we would be able to identify periodic attentional reorientations during early as well as later development (17,47); later in development, however, we predicted that infants would be more likely to extend visual fixations beyond their modal periodic reorientation rate, possibly indicating a greater or more efficient integration of attention and gaze shifting (48), and that attention duration episodes would be longer overall.

Next, in order to assess the link between lower-level mechanisms of autonomic arousal and attention, we calculated cross-correlations between autonomic arousal (indexed via heart rate) and attention episodes across the entire play session for both 5- and 10-months olds. This allowed us to examine whether arousal changes tend to forward-predict changes in attention, or vice versa (analysis 2). Based on previous research (49), we hypothesized that periods of elevated autonomic arousal would associate with, and forwardpredict, shorter attention episodes. We also predicted that such relationship would weaken with time due to the maturation of cortical attentional systems.

We used a similar approach to examine developmental changes in the relationship between neural markers 125 126 of executive attention and real-world attention behaviours. We were interested to examine whether neural changes (indexed by theta power) anticipate subsequent attentional behaviour shifts (48); or, whether neural 127 processes after the attention shift relate to increases in infants' attention engagement. To test this, we 128 129 conducted three analyses. First, we analysed neural activity across a range of time windows both before and after the onsets of new attention episodes and performed linear mixed effect models to examine how 130 131 neural activity before and after attention onset associated with the subsequent durations of those episodes 132 (analysis 3). Second, we examined changes in neural activity during individual attention episodes (analysis 4). Finally, we used cross-correlations to examine whether, across the entire dataset, neural markers tend to 133 134 forward-predict changes in attention, or vice versa (analysis 5). We predicted that the associations between 135 neural markers of executive attention and real-world attention behaviours would become stronger with increasing age (i.e., theta activity would show a stronger predictive relation with infants' attentional 136 137 behaviours at 10 months, as evidence of increased modulatory power from the executive attention system 138 on infants' attention).

Finally, we examined whether there were any interdependencies between autonomic arousal and theta activity. We had no predictions for how this relationship would change over time.

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- 142 **2. Results**
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144 2.1 Analysis 1: Developmental changes in attention

Our first set of analyses examined attentional inertia as a measure of internally driven attentional engagement. Attention inertia is the phenomenon that, as individuals become progressively more engaged with an object, their attention progressively increases. We tested whether attention inertia is stronger at 10 months compared to 5 months. To do so, we calculated the ACF and the survival probability of spontaneously occurring attention episodes during play to quantify the rate of change of spontaneous attention durations and the probability between looking (i.e., paying attention) and looking away, respectively.

152 Initially, we conducted four descriptive analyses to test how attention and inattention durations and reorientations change over both the course of the solo play interactions and developmental time. First, we 153 154 tested how many times per minute 5- and 10-months-old infants redirected their attention from one object to the other. We found that, on average, 5-month-old infants performed significantly more both attentive 155 (t(10)=4.346, p=0.001) and inattentive (t(10)=4.202, p=0.002) reorientations during the solo play 156 157 interaction than 10-months-old infants (Figure 2A). When we looked at how attention reorientations changed during the course of the solo play episode, we found that 5-month-old infants performed 158 159 consistently more looks than 10-month-olds throughout the interaction even though the number of looks 160 per minute decreased over the course of the interaction for both age groups (Figure 2B, and Figure S1A and B). 161

162 Second, we investigated the average duration that 5- and 10-month-old infants spent in attentive and in inattentive states during the solo play interaction and minute by minute (Figure 2C and D respectively). In 163 164 general, infants' attention durations toward play objects at 10 months were longer (t(10)= -2.787, p=0.019). At 5 months, moments of inattention were longer than moments spent looking towards the object (t(10) =165 166 -3.749, p=0.003). Overall, at 10 months, but not 5 months, infants spent more time in attentive compared 167 to inattentive states (t(58)=10, p<0.001) (Figure 2E). We then calculated a best fit line, individual by individual, to look at how average attention duration changed within the session (see Figure S1C and D). 168 We found no significant differences in the way attention duration changed during the interaction between 169 170 the two age groups (Figure S1C and D).

Third, we explored the distribution of looks towards the objects (Figure 2F). At both ages, attention durations shorter than or equal to 5 seconds follow a positively skewed lognormal distribution, with modal 175 episodes. There was an increasing amount of such looks with increasing age.

Following the descriptive statistics, we calculated both the ACF and the survival probability of the looking 176 177 behaviour (Figure 3). First, we used time-series analyses to examine the rate of change of attention durations, relative to itself. We calculated the ACF of the attention durations at both time points (more 178 179 details in 4.3.5.2). The ACF indexes the cross-correlation of a measure with itself at different lag-intervals in time (46). ACF values were obtained from 0 to 10 seconds lag, in steps of 500 milliseconds. As shown 180 181 in Figure 3A, the ACF of the time series looking behaviour fell off more sharply at 5 months than at 10 months. The ACF values were compared across ages using independent sample t-tests. From lag +500 182 183 milliseconds to 10 seconds, 10-months-old infants showed significantly higher correlation values than 5-184 months-old infants.

185 Second, we performed a survival analysis by calculating the survival probability function of the looking behaviour towards the objects at both time points. The survival probability function is the probability that 186 an attention episode survives longer than a certain time. As shown in Figure 3B, the survival probability of 187 188 a look decreased abruptly at the beginning, for the very short looks, and flattened as looks got longer. The 189 differences in the speed at which the survival probability decreased can be seen more clearly by calculating the derivative of the survival probability (Figure 3C). To compare survival among the two groups, we 190 191 performed the log rank test using the Matlab function 'Logrank' (50). The results for the log-rank test 192 rejected the null hypothesis (p<0.001) indicating that the survival curves for looking behaviour at 5 months 193 and 10 months were significantly different. Notably, the likelihood of a look ending is more tightly clustered 194 around the modal value of 0.5 seconds at 5 months.

Overall, our results showed that older infants demonstrated to have a slower-changing profile of attention with longer attention episodes overall (Figure 2A and B, Figure 3A). At both ages, there was evidence for a preferred modal reorientation rate in the 0.5-0.6 second range, which was slightly faster at 5 months than 10 months (Figure 2F). Attention durations were more tightly clustered around the modal value at 5 months. At 10 months, attention episodes were more likely to be extended beyond the preferred modal reorientation rate than at 5 months (Figure 3B and C).

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2022.2Analysis 2. Auto- and cross-correlation analyses between infant autonomic arousal and203attention

In this section we investigated the relationship between changes in infant autonomic arousal (indexed by heart rate activity) and their associations with moment-to-moment changes in attention (indexed as a continuum of looking durations to play objects vs. elsewhere).

Figure 4 (A and C) shows the results of the autocorrelation analyses for autonomic arousal at 5 and 10 207 208 months of age respectively. Significant autocorrelations were observed at relatively short lags around t=0 209 (from -4 to +4s) at both ages. Figure 4 (B and D) shows the results of the cross-correlation analysis between 210 autonomic arousal and attention at 5 and 10 months of age respectively. The negative values indicate that, at 5 months, lower heart rate forward-predicted increased looking durations from lags between -9 to -2 211 212 seconds (i.e., lower heart rate at time t significantly associated with increased attention at time t+9 seconds). 213 The same pattern was present but not significant at 10 months. The asymmetry of this cluster around the 214 lag t=0 indicates that changes in heart rate tended to forward-predict changes in attention more than vice 215 versa.

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217 **2.3** Analysis 3. Calculation of neural power changes around an attention episode

We used linear mixed effects models to examine the associations between the length of each attention episode (i.e., looking duration to any of the play objects) and relative theta power at different time windows relative to the onset of that attention episode (see Figure 5). At 10 months, relative theta power in the time window of 0 to +1000msec and +1000 to +2000msec after onset of a new attention episode predicted the subsequent duration of that attention episode. At 5 months, the same relationships were not significant. We found no evidence of neural activity before the start of an attention episode forward-predicting the length of that attention episode at any time point (Figure 5).

225 The final number of accepted trials (i.e., attention episodes) in the analyses varied across the three timewindows immediately after the onset of each look. More trials were obtained for the first window (total 226 227 number of looks at 5 months was 790, and 411 at 10 months) than the second (total number of looks at 5 228 months was 473, and 336 at 10 months) and the third (total number of looks at 5 months was 301, and 277 229 at 10 months). All three conditions ended up with enough number of clean trials that was greater than the 230 recommended number of trials in the infant EEG literature (51-53). Thus, the differences between the 231 number of trials for each time window are not expected to contribute to the results described above. However, we repeated this analysis by matching the number of attention episodes at 5 months to the ones 232 analysed at 12 months. We found no differences in the results (see Figure S2). 233

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235 2.4 Analysis 4. Calculation of neural power changes within an attention episode

In addition to the previous analyses, which examined the associations between the length of each attention 236 237 episode and relative theta power at different time windows relative to the onset of that attention episode, we also wished to examine whether power at the theta and alpha band changed significantly during an 238 239 attentional look (i.e., any look at a play object) (Figure 6). Relative theta was analysed as a function of these three factors: time within an attentional episode, brain areas and age with a 3-way ANOVA (Figure 6). 240 There was no statistically significant interaction between the three factors. However, the analysis revealed 241 two simple two-way interactions: one between time within an attention episode and age, F(2) = 5.58, p < 242 243 .005 and the other between channel cluster and age F(2) = 11.98, p < .001. Next, we performed a multiple 244 comparison test to find out which groups of factors were significantly different. Results are shown in table S1-S3. A follow up analysis showed a significant effect of "time within an attentional episode": 10-months-245 old infants had greater theta during the third-to-fourth second into the look (middle) than the first second 246 (start) in both the central and the frontal poles. These effects were not present in 5-months-old infants. 247 Similarly, relative alpha was also analysed as a function of these three factors: time within and attentional, 248 249 brain areas and age with a 3-way ANOVA. We found no statistically significant interactions. Results are shown in Figure S3. 250

Again, the final number of accepted trials (i.e., attention episodes) in the analyses varied across the three time-windows into each look. More trials were obtained for the first-second window (total number of looks at 5 months was 791, and 415 at 10 months) than the third-to-fourth second (total number of looks was 172 at 5 months, and 194 at 10 months) and the last second before look termination (total number of looks was 476 at 5 months, and 337 at 10 months).

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257 **2.5** Analysis 5. Auto- and cross-correlation analyses between infant theta activity and attention

In this section we investigated the relationship between dynamic changes in infant endogenous brain activity and their associations with moment-to-moment changes in attention (measured as a continuum of looking time durations to play objects vs. elsewhere). Figure 7A and D shows the results of the autocorrelation analyses for infant theta activity. Figure 7B and E shows the results for the cross-correlation analyses between infant theta activity and infant attention. Cluster-based permutation analysis revealed a significant positive association between the two variables (marked with a red line) at 10 months around time lag=0. More specifically, increases in infant theta activity at 10 months were significantly correlated with fluctuations in infant attention (Figure 7E). No associations were found between theta activity and infant attention at 5 months of age. Interpreting the exact time intervals over which a cross-correlation is significant is challenging due to the autocorrelation in the data (54,55), but the fact that the significance window is asymmetric around time 0 indicates a temporally specific relationship between infant attention and theta power, such that attention forward-predicts theta power more than vice versa.

Finally, to test whether there were any interdependencies between autonomic arousal and brain activity, we performed a cross-correlation analysis between these two variables. We found a significant cluster at 10 months (Figure 7F) but not at 5 months (Figure 7C). The asymmetry of this cluster around t=0 indicated that changes in brain activity tended to precede changes in autonomic arousal more than vice versa.

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276 **3. Discussion**

We examined developmental changes in the physiological and neural correlates of real-world attention patterns during early development. To do so, we measured attention durations (to an accuracy of 50Hz), along with cortical neural activity (EEG) and autonomic arousal (via ECG) from typical 5- and 10-monthold infants playing alone while seated at a tabletop with 3 toys. This age range is a key period for early cognitive development, as differential patterns of brain development (10) drive a transition from primarily subcortical to cortical control (56), and early-emerging atypicalities can have life-long consequences (57,58). However, many of the mechanisms that drive early development remain unclear.

284 From Analysis 1 we found that infants at both ages showed a preferred rate of reorientation (i.e. their visual attention took place periodically). The modal durations of attention episodes towards different play objects 285 286 were in the 0.5-0.6 second range at both ages but were lower at 5 months (Figure 2F, 2C). This contrasts with analyses of micro-level fixation durations (time intervals between individual refoveating eye 287 288 movements), which decreases from early infancy (~ 0.5 secs) through to later infancy (~ 0.4 secs) through to adulthood (~ 0.3 secs) (15,18). Research with adults suggests that the minimum time necessary to plan 289 290 and executive a saccade is ~80msecs in adults (16). Although the equivalent figure is not known in infancy, the fact that modal attention durations towards objects were shorter at 5 months than 10 months, whereas 291 292 fixation durations decrease with age, makes it likely that the figures we observed do not simply indicate 293 that infants were reorienting at the fastest speed possible, but rather were reorienting according to a 294 preferred modal reorientation rate (15).

295 The survival analysis showed that, at both ages, looks were fragile early in their existence and most likely 296 to terminate in the <1 second range (45) but the speed at which the survival probability curve decreased 297 was faster at 5 months, meaning that the probability of a look lasting longer than time t was lower at 5 months. Richards and colleagues have found similar relationships in infants in both lab-based and 298 naturalistic settings (45). Overall, attention durations were shorter at 5 months; this faster-changing pattern 299 300 of attention to the object was also reflected in the ACF of their looking behaviour, which decreased significantly faster, showing lower overall self-similarity. Collectively, these data fit well with what we 301 302 know about the development of attention. With time, we seem to observe a higher-level control of attention that allowed older infants to prioritize the task at hand – learning about/ exploring the toys – as well as to 303 304 inhibit the tendency to shift attention away from an interesting task (6,20,25). Alternatively, longer attention 305 episodes might arise because children physically manipulate objects, bringing objects closer to themselves which makes them more exogenously salient (37,59). In this case, then the infant's increased looking 306 behaviour would be the result of increased exogenous attentional capture rather than an increase in 307 endogenous attention control (46). 308

309 Analysis 2 examined how dynamic fluctuations in autonomic arousal relate to moment-to-moment changes 310 in attention. Consistent with previous work (13,49,60), the average concurrent correlation between autonomic arousal and attention was negative at both age points, indicating that lower arousal was 311 associated with increased likelihood of attention. Such links have been considered within the developmental 312 attention regulation literature, where increases in arousal are thought to lead to distraction or difficulties 313 focusing attention, and vice versa (61). We also found that arousal levels were significantly forward 314 predictive of attention at 5 months but not at 10 months (Figure 4B and D). Theoretically, if attentional 315 episodes drive decelerations in the heart rate (12), and older infants show longer attentional episodes on 316 average, then one could hypothesise that older infants ought also to show a more stable pattern (i.e., higher 317 autocorrelations) in their heart rate fluctuations than younger infants. However, this was not what we 318 319 observed (Figure 4A and C). Overall, the much shorter attention durations observed in this setting, compared with screen-based TV viewing (12), means that heart rate decelerations relative to individual 320 321 attention episodes were observed infrequently in our data. However, our data did suggest, consistent with 322 previous research, that at 5 months, changes in autonomic arousal forward-predict subsequent changes in 323 attention.

In Analysis 3 we examined the associations between attention episode durations and theta power either before, or after, onset of that attention episode. At 10 but not 5 months, increased theta during the period immediately after the onset of a new attention episode (0-2000msec) forward-predicted the subsequent length of that attention episode (Figure 5). At neither age, however, did cortical neural activity before theonset of an attention episode forward-predicted attention durations.

In Analysis 4 we examined whether cortical neural activity changed significantly during an attention episode. Consistent with previous research (29), theta power in central and frontal electrodes increased significantly during an attention episode at 10 months, but not at 5 months (Figure 6). Contrary to our expectations, we did not find a link between attenuated alpha during an attention episode at any age (see Figure S3).

In Analysis 5 we investigated the relationship between dynamic changes in infants' theta activity and moment-to-moment changes in attention. We identified a significant positive association between infant theta activity and infant attention at 10 months but not at 5 months (Figure 7B and E). Interpreting the exact time intervals over which a cross-correlation is significant is challenging (54,55), but the asymmetry of the cluster around time 0 indicates that attention forward-predicted theta power more than vice versa, consistent with the findings from Analysis 3.

These findings are consistent with previous research suggesting that, by 10 months, but not during early infancy, theta oscillations increase during sustained attention and encoding (29,41,62) and associate with longer attentional periods (63). Importantly, though, we found no evidence that endogenous neural markers before the onset of an attention episode forward-predict the length of an attentional episode at either age. Instead, what we found suggests that neural activity shortly after the onset of an attention episode forwardpredicts the length of that episode. One possible interpretation of this is that neural activity associates with the maintenance more than the initiation of attentional behaviours (64).

347 Finally, we examined the relationship between theta power and autonomic arousal (Figure 7C and F). A cross-correlation analysis found a negative forward-predictive relationship between the two, such that 348 349 increases in theta forward-predicted decreases in autonomic arousal at 10 months, but not at 5 months. This suggests that changes in the brain activity could be modulating subcortical changes (i.e., changes in the 350 351 heart rate) and may thus be able to initiate or maintain states of arousal that are common to vigilant or 352 sustained attentional states (6,12). Overall, it appears that, by 10 months, the different substrates of attention 353 are more inter-linked, and stronger associations are emerging between behaviour, cortical activity, and 354 autonomic arousal (65).

In summary, our results suggest that, earlier in development, attentional episodes are more influenced by lower-order endogenous factors such as a general arousal system (3,12) - that might reflect a stronger influence of subcortical structures over the modulation of attention - and show a preferred modal 358 reorientation timer - which characterises infants' attention shifting more strongly. Such factors would also 359 be present at older ages; however, their association with attention would weaken over developmental time 360 due to the maturation in cortical attentional areas thought to take place throughout the first year of life. Later in infancy, cortical neural activity reliably changes during attention episodes, but does not forward-361 predict attention at either age; rather, it seems that neural changes associate with the maintenance more than 362 the initiation of attentional behaviours. Overall, the modulation of attention seems to involve both arousal-363 based and cortical processes. With developmental time, however, the latter increases its control over the 364 modulation of both (i.e., overt attentional behaviours and arousal), resulting in a more inter-linked system 365 where associations between attentional systems are stronger. Theoretically, this is consistent with what we 366 know about the development of executive attention from experimental and neuroanatomical studies. 367

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369 **3.1 Limitations and strengths**

370 Our findings should be interpreted with consideration to a number of limitations of the study. First, our 371 events of interest are intrinsically linked with one of the biggest EEG artefacts (i.e., eve movements), and 372 so it is possible that residual artifact in the EEG signal may have contaminated our data. However, our data 373 were processed using algorithms specially designed to clean naturalistic EEG data (66,67), and previous analyses suggest that the electrode locations and frequency bands that we examined should be least affected 374 375 by artifact, compared with more anterior locations and higher and lower frequencies (68). Additionally, our 376 analyses were carefully designed to preclude this potential confound. First, our analyses compare events that we know share the same level of artefact/ noise (i.e., saccades at 5 months old contribute to comparable 377 378 noise levels than at 12 months old (69); second, analysis 3 and 4 are time-locked to a saccade to eliminate the possibility that saccadic frequency may have influenced our results; and third, other research (70) 379 380 suggest that artifact associated with saccades disappears within ~300msecs, whereas the associations 381 between theta and look duration lasts much longer than this, up to ~6 seconds. For all this, we consider that the possibility that our results may have been caused by infants' saccades is unlikely. 382

Second, while the study of theta and alpha activity can offer insights into infants' intrinsically guided attention beyond its behavioural manifestations (26), attributing functional significances to particular frequency bands can be risky, especially when the ways in which they are controlled and the extent to which they interact across attention-related brain networks, still remain largely unknown (71). Similarly, the fact that certain frequency bands seem to covary with attention does not exclude the possibility of their correlation with other processes as well. Third, the use of different EEG systems (32- vs. 64-channel BioSemi gel-based ActiveTwo) and age groups might have contributed to the differences we observed over time. However, we compared the EEG signal quality between groups and found no significant differences (Table S5-6, Figure S4).

Fourth, we used a different set of toys at the two ages (see Figure S5). Consequently, this introduced a new source of variation (i.e., toy characteristics) that might have contributed to any of the observed differences (20). However, we chose to present developmentally appropriate stimuli at the two ages to ensure that the cognitive demands were similar at the two ages. Thus, while still possible, it is unlikely that the developmental differences observed in the current study might be due to differences in the amount of information processing on the part of the infant and/ or the "interestingness" of the toys.

Fifth, it is worth mentioning that, while infants gather information about their world through aggressive visual foraging, looking and attending are not synonymous. Previous research has shown that covert shifts of attention can occur without overt shifts of gaze by 4-6 months of age (72,73). However, the current study has focused exclusively on overt attention.

Finally, our laboratory setting was a novel environment for our participants and might have elicited behaviours that are different from the ones that develop at home. However, it still represents a significant advancement relative to other screen-based or highly controlled experimental tasks.

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407 **4. Materials and Methods**

408 4.1 Experimental Design

Looking behaviour, EEG and ECG data were collected from mothers and their infants at two age points: 5 409 and 10 months while playing alone. At 5 months, infants were seated either in a highchair or on a 410 researcher's lap and a table was positioned in front so that toys on the table were within easy reach (see 411 Figure 1). To reduce infant's stress, mothers were present in the room but moved to another smaller table 412 on the right side of the original table and given an identical set of toys which they played with in parallel. 413 A wooden divider was positioned between the two tables to prevent infants from seeing the objects with 414 which their mothers were playing. At 10 months, the same procedure was used but the divider was 415 416 positioned across the midline of the table and the adult participants were seated directly opposite the infants. In both situations, mothers and infants had direct line of sight to one another but neither could see the 417 418 others' toys on the table.

419 The same three age-appropriate toys were always used for each age group. These were small and relatively 420 engaging (see Figure S5). During the solo play interaction, one of the researchers sat behind the infant to 421 collect the toys that fell on the floor (either because the infants threw them or because they fell from their hands) and brought them back on the table. Mothers were allowed to speak during the interaction but were 422 instructed not to name the toys they were playing with to prevent infants from the influence of any 423 exogenous parental' influence. In the Supplementary Materials, we present a set of analyses that preclude 424 the possibility of maternal influence on infants' behaviour and demonstrate that the impact of the mothers 425 on the infants' behaviour did not differ between age groups (SI 3). The average duration of the interactions 426 with usable EEG/ ECG data did not differ significantly between 5 and 10 months (interactions with EEG 427 (average duration at 5m = 292.4s, and 10m = 250.1s, t(46)= -1.85, p= 0.07); interactions with ECG data 428 (average duration at 5m = 351.2s, and 10m = 317.9s, t(40) = -1.1, p = 0.27)). 429

The interactions were filmed using three Canon LEGRIA HF R806 camcorders recording at 50 fps. At 5 months, one camera was placed in front of the infant and another one was placed in front of the mother. At 10 months, two cameras faced the infant: one placed on the left of the divider, and one on the right. The other camera faced the mother and was positioned just behind the right side of the divider. All cameras were placed so that the infant's and the mother's gaze, as well as the three toys placed on the table, were always visible.

Brain activity was recorded using a 64-channel at 5 months and a 32-channel at 10 months, BioSemi gelbased ActiveTwo system with a sampling rate of 512Hz with no online filtering using Actiview Software.

Heart rate activity was recorded using a BioPac[™] (Santa Barbara, CA) system recording at 2000Hz. ECG
was recorded using disposable Ag-Cl electrodes placed in a lead II position.

440

441 4.2 Participants

Participants were typically developing infants and their mothers. The catchment area for this study was East London, including boroughs such as Tower Hamlets, Hackney and Newham. Participants were recruited postnatally through advertisements at local baby groups and local preschools/ nurseries. We also operated a word-of-mouth approach, asking parents who got involved to ask if their local networks would be interested in participating. Ethical approval was obtained from the University of East London ethics committee (application ID: ETH2021-0076). Informed consent, and consent to publish, was obtained by the caregivers of the infants tested. Initial exclusion criteria included complex medical conditions (e.g., heart rate condition, neurological/ genetic abnormality), known developmental delays, prematurity, uncorrected vision difficulties, and parents below 18 years of age. Further exclusion criteria as well as final numbers of data included in each of the analyses for both samples are summarised in Table S7. The final sample included 12 infant females and 19 infant males at 5 months and 14 infant females and 15 infant males at 10 months. Data was analysed in a cross-sectional manner. Average age was 5.32 months (std = 0.58) and 10.49 months (std = 0.87). This is the first time that any of this data has been analysed and reported.

Since the analyses are performed conducted relative to specific events (such as the frequency and duration of looks to objects), and each participant averaged around 10 looks per minute at 5 months and approximately 4 looks at 10 months (see Figure 1A and B), we believe that the relatively low N for this study is balanced by the considerable amount of data points accessible.

460

461 **4.3 Data processing and Statistical Analysis**

462 4.3.1 Synchronisation between behavioural and EEG/ ECG data

The cameras were synchronised to the EEG and ECG via radio frequency (RF) receiver LED boxes attached to each camera. The RF boxes received trigger signals from a single source (computer running Matlab) at the beginning and end of the play session, and concurrently emitted light impulses, visible from each camera. At the same time, triggers were sent and stored in the Actiview Software and recorded to the EEG data as well as to the Acknowledge Software and recorded to the ECG data.

The video coding and EEG/ ECG data synchronisation was done by aligning the times of the LED lights and the EEG/ ECG triggers. We also checked for dropped/missing frames by checking that the time between the LED lights matched the times between the EEG/ ECG triggers.

471

472 **4.3.2 Video coding**

The looking behaviour of the infants was manually coded offline on a frame-by-frame basis, at 50fps. The start of a look was considered to be the first frame in which the gaze was static after moving to a new location. The following categories of gaze were coded: looks to objects (where the infant was focussing on one of the three objects), looks to partner (where the infant was looking at their partner), inattentive (where 477 the infant was not looking to any of the objects nor the partner) and uncodable. Uncodable moments 478 included periods where: 1) the infant's gaze was blocked or obscured by an object and/or their own hands, 479 2) their eyes were outside the camera frame, and/ or 3) a researcher was within the camera frame and the infant turned to them and/or realised a researcher was around. Video coding was completed by three coders, 480 who were trained by the first author. To assess inter-rater reliability, ~15% of our data (10 datasets) were 481 double-coded by a second coder and Cohen's kappa was calculated. There was moderate agreement ($\kappa =$ 482 0.581, std=0.183) (74). Due to the unusual nature of our behavioural coding (with gaze coded across many 483 484 20ms bins) the interrater reliability is heavily contingent on how we calculate it. We chose to report the most stringent calculator of inter-rater reliability. 485

Looking behaviour data was then processed such that any look preceding and following an "uncodable" period was NaN-ed and excluded from further analyses. Similarly, both the first and the last look of every interaction were also NaN-ed and excluded from further analyses.

489

490 **4.3.3 EEG artefact rejection and pre-processing**

491 EEG data was pre-processed and cleaned from oculomotor and other contaminatory artefacts using a fully 492 automatic artefact rejection procedure specially designed for naturalistic infant EEG data by Mariott Haresign (66), building on previous related work (59, 60). Briefly, this involved the following steps: 1) data 493 494 were high-pass filtered at 1Hz, 2) line noise at 50Hz was eliminated using the EEGLAB function clean line.m, 3) data were low-pass filtered at 20Hz, 4) the data were referenced to a robust average 495 496 reference 5) noisy channels were rejected using the EEGLAB function pop_rejchan.m, 6) the channels 497 identified in the previous stage were then interpolated back, using the EEGLAB function eeg_interp.m, 7) continuous data were automatically rejected (NaN-ed) in a sliding 1s epoch based on a percentage of bad 498 499 channels (set here at 70% of channels) that exceed 5 standard deviations of the mean channel EEG power, and 8) Independent Component Analyses (ICA) were computed on the continuous data using the EEGLAB 500 501 function runica.m. Only participants with fewer than 30% of channels interpolated at 5 months and 25% at 10 months (step 6) made it to the final step (step 8, ICA) and final analyses. To compare the quality of the 502 503 EEG data at 5 and 10 months we performed a series of analyses on percentage of channels interpolated, total segments removed (i.e., zeroed out) and total ICA components rejected (see Table S5-6 and Figure 504 505 S4).

The higher density array was down sampled so that all the EEG analyses described below used only shared channels between the 32- and the 64- channel EEG systems. We selected three main clusters of electrodes for our analyses: Frontal channels ('Fp1', 'Fp2', 'AF3', 'AF4', 'Fz'), Central channels ('FC1', 'FC2', 'C3', 'Cz',
'C4', 'CP1', 'CP2'), and Occipital channels ('PO3', PO4', O1', 'Oz', 'O1') (see Figure S6).

510

511 **4.3.4 Heart Rate – Beats per minute**

512 R-peak identification was performed using the in-built Matlab function 'findpeaks'. The minimum peak height was manually defined as a simple amplitude threshold after visualising the raw data, minimum peak 513 514 distance, instead, was set at 230msec. Following this, automatic artefact rejection was performed by 515 excluding those beats showing an inter-beat interval (IBI) <330 or >750msec (i.e., allowing a minimum of \sim 80BPM and maximum of \sim 180BPM), and by excluding those samples showing a rate of change of IBI 516 517 greater than 90msec between samples. Next, we converted IBI values into beats-per-minute (BPM) values and removed outliers in the BPM time series. These were defined as values falling 2.5 interquartile ranges 518 519 above the upper quartile and below the lower quartile. Outliers were then interpolated using the Matlab inbuilt function 'fillmissing' with the 'spline' method. Finally, we epoched the data into one-second epochs 520 by averaging all the BPM values comprised in each one-second epoch. 521

522

523 **4.3.5** Analysis 1. Developmental changes in attention

524 4.3.5.1 Overt attention and inattention extraction

The aim of our analysis was to identify moments where the infants paid attention to (i.e. looked at) any of the play objects as opposed to inattentive moments. Accordingly, all looks to object and inattentive looks were selected and categorised as attentional and inattentive episodes, respectively. Looks to partner were excluded from all analyses.

Following this, we extracted the first and last frame of all looks of interest (i.e., looks to objects and inattentive looks). To calculate attention and inattention durations, we subtracted the last frame from the first frame of each look of interest and divided it by the sampling rate (i.e., 50) to convert "duration in frames" to "duration in seconds". Attentive and inattentive reorientations (count) were calculated by counting the occurrence of each of these two attentional states and dividing it by the length of the session.

534 4.3.5.2 ACF of the attention duration

Here, we extracted the duration (in seconds) of all the attentional episodes that happened within the play session and zero-ed out all the non-attentional episodes. This allowed us to create a time series string with the durations of each consecutive attention episode. We then calculate the autocorrelation of that signal and repeated these steps for each behavioural dataset. Finally, we averaged the ACF values within each age

- 539 group to obtain the ACF values reported in Figure 3A.
- 540

4.3.6 Analysis 2. Auto- and cross-correlation analyses between infant autonomic arousal and attention

543 4.3.6.1 Attention, one-second epochs

To calculate the "attention" variable, we epoched the gaze data into one-second epochs and calculated the duration (in msec) of each attention episode relative to the one-second epochs. Most epochs were coded as either 1000 (epochs where the child was attending throughout) or 0 (epochs where the child was inattentive throughout). If an attention episode started halfway through a one-second epoch, then it was coded as 500. The other non-attention episodes (i.e., inattentive looks, looks to partner) were zeroed. See Figure 1 for a schematic view of the procedure we followed to parse the looking behaviour into one-second epochs.

550 4.3.6.2 Cross-correlation analyses

To investigate the relationship between autonomic arousal and fluctuations in attention we performed a cross-correlation analysis between the two variables. Importantly, these analyses are not time locked to specific moments (i.e., start of an attentional event) and are conducted on two time series (i.e., attention and heart rate fluctuations) as a whole. Because of this, the strength of the overall correlation is weakened by the fact that periods of expected stronger correlation are balanced by weaker correlations where we would not necessarily expect any correlation at all.

Additionally, we also computed the autocorrelation for autonomic arousal to assess how well it predicts itself over time and evaluate its stability. All analyses were computed at lags from -30 to +30s in 1s intervals. The cross-correlations values at each time-lag were computed individually and then averaged across all participants. The procedure was identical for the autocorrelation, except that instead of examining the relationship of two different time series at variable time intervals, we assessed the relationship of one time series to itself at variable time intervals.

563 To assess significance of the cross-correlations, we first used bootstrapping to generate confidence 564 intervals, using an approach that controls for the level of autocorrelation in the data. To do this, the time series data of one participant (e.g., attention of participant 1) was randomly paired with the time series data 565 of another participant (e.g., autonomic arousal of participant 13). If the time series datasets had different 566 lengths (due to different participants having different session lengths), we appended zeros to the end of the 567 shorter vector to match the length of the longest vector. We then computed the cross-correlation between 568 all the unique combinations that could be found within each sample (e.g., in a sample N=23, the maximum 569 of unique combinations is 529). Next, the cross-correlation results in the permuted data were randomly 570 grouped in samples that were the same size as the original data (e.g., N= 23) and averaged together. This 571 procedure was repeated 1000 times and used to generate the 95% confidence intervals. In this way, we 572 573 identified whether the observed cross-correlation values at each time interval differed significantly from 574 chance.

575 Next, to control for multiple comparisons across time intervals, we used a cluster-based permutation 576 approach. On each iteration, one permutation was compared with the 999 other permutations, and 577 significant time-points were identified as values falling above the 97.5th centile and below the 2.5th centile 578 (corresponding to a significance level of 0.05). We then identified the two largest clusters of significance that occurred by chance: one for positive correlation values and the other for negative correlation values. 579 We repeated this 1000 times. Following this, we created a distribution of cluster sizes for positive and 580 negative correlation values and took the size value corresponding to the 95th percentile in each distribution 581 to define our cluster-size threshold. Finally, we compared the cluster sizes obtained in the observed data 582 against the cluster-size threshold and only considered significant the ones that exceeded such threshold. 583

584 Calculating the significance levels of the autocorrelation was more straightforward. This was done by first 585 calculating the autocorrelations based on individual datasets, and then averaging the significance values of 586 the Spearman's correlations at each time interval.

587

588 4.3.7 Analysis 3. Calculation of neural power changes around an attention episode

We examined the associations between the duration of infant attention episodes and infant theta changes around these looks using linear-mixed effects models. Infant attention episodes and attention duration were calculated as explained above in section 4.3.5.1. 592 To conduct these analyses, each infant attention episode onset (i.e., gaze shift to a different toy) was 593 identified in the EEG signal by calculating the time from the start of the interaction (first LED) to the onset 594 of the look (in the behavioural data) and adding it to the first EEG trigger. For each look, we extracted theta (3-6Hz) power for two time-windows immediately prior to the onset of each look (-2000 to -1000msec and 595 -1000 to 0 msec pre-look onset) and three time-windows immediately after the onset of each look (0 to 596 597 1000ms, 1000ms to 2000msecand 2000 to 3000msec post look onset) (see Figure 1). To calculate the EEG power spectra, we use the 'mtmfft' method from the ft freqanalysis function in FieldTrip, an open-source 598 599 Matlab toolbox (75); http://fieldtriptoolbox.org). Extreme power values that were 4 times greater than the interquartile range were treated as outliers and excluded from further analyses (similar to 29). More detail 600 601 on the amount of data available (i.e., average duration of the session per participant and number and duration 602 of attentional episodes per minute) can be accessed in Table S1.

For each epoch, we only selected power within our cluster of central channels (similar to 60). Power at each bin was expressed as relative power, defined as the total power at a specific frequency band (e.g. 3 to 6Hz for theta) divided by the total power across all frequency bands (1 to 20Hz) during that epoch. After extracting the relative power in the theta band, we calculated separate linear mixed effects models for each of the five windows to examine the relationship between EEG power within that time window and attention duration.

609

610 **4.3.8** Analysis 4. Calculation of neural power changes within an attention episode

In addition, we also wanted to look at power changes within attention episodes. Infant attention episodes 611 612 and attention duration were calculated as explained above (section 4.3.5.1) and each infant look onset towards an object was identified in the EEG signal as described in analysis 3 (section 4.3.7). For each look, 613 614 we extracted the first (0 to +1000msec, "start") and third-to-fourth (3000 to 4000msec, "middle") second into the look, and the last second (-1000msec prior to look offset to look offset, "termination") before look 615 616 termination (see Figure 1). Looks that did not make it to the full second segment were excluded from further analysis. Similarly, only looks that were longer than 5 seconds were included to the "middle" group. This 617 618 was done to avoid an overlap between the activity from the "middle" and the "termination" groups.

Frequency analysis was conducted to assess the power spectral density for both theta (3-6Hz) and alpha (6-9Hz) frequency rhythms for each of the three time-segments. These analyses were calculated for the three prespecified clusters of channels: Frontal, Central and Occipital (see Figure S6). Again, power at each time segment was expressed as relative power. The selection of both theta (3-6Hz) and alpha (6-9Hz) frequency bands was led by previous work using this same approach (e.g., 28, 29, 30, 37).

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626 **4.3.9** Analysis 5. Auto- and cross-correlations analyses between infant theta activity and attention

627 4.3.9.1 EEG relative power, one-second epochs

For this analysis, we parsed the EEG data into one-second segments and calculated the relative theta powerfor each one-second segment as described above (see 4.3.7).

630 4.3.9.2 Cross-correlation analysis

To explore whether modulations in endogenous theta activity related to fluctuations in infants' attention, we conducted a cross-correlation analysis between infants' relative theta and attention. Attention was calculated as described in analysis 2 (4.3.6.1, and Figure 1). Additionally, we also computed the autocorrelation for relative theta to assess how theta predicts itself over time. Again, all analyses were computed at lags from -30 to +30s in 1s intervals. Significance was assessed following the steps described in analysis 2 (4.3.6.2).

Finally, to explore interdependencies between autonomic arousal and theta activity we conducted a cross-correlation analysis between infants' autonomic activity and relative theta.

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642 **References**

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814 Human subjects: Ethical approval was obtained from the University of East London ethics committee

815 (application ID: ETH2021-0076). Informed consent, and consent to publish, was obtained by the caregivers

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- 817 in Figure 1.
- 818

819 Author contributions:

- 820 Conceptualization: SVW, EJHJ, MPA.
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831

832	Competing interests:
833	Authors declare that they have no competing interests.
834	
835	Data and materials availability:
836	Partial restrictions to the data and/or materials apply. Due to the personally identifiable nature of this data
837	(video recordings from infants) the raw data will not be made publicly accessible. Researchers who wish
838	to access the raw data should email the lead author. Permission to access the raw data will be granted as
839	long as the applicant can guarantee that certain privacy guidelines (e.g. storing the data only on secure,
840	encrypted servers, and a guarantee not to share it with anyone else) can be provided.
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846 Figures and Tables

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Figure 1. Experimental set up and schematic illustration of the procedure followed for analysis 2 to
5. Top figure shows the experimental set up for solo play. Below, on the left, is shown the procedure

followed to parse the looking behaviour and create the variable "attention", and further cross-correlation
analyses (analysis 2 and 5). On the right, instead, is shown the steps followed to identify attentional episodes

- in the EEG signal and further EEG analyses (analysis 3 and 4).
- 853

Figure 2. Descriptive analyses on infant attentional behaviour during the solo play interaction. (A) 854 Average number of attentive and inattentive looks per minute at 5 months (left) and 10 months (right), (B) 855 856 Average number of attentive and inattentive looks minute by minute. Asterisks show the significance values of comparisons examining how average number of looks per minute of the interaction differed as a function 857 of age. (C) Average duration spent in one of the two possible attentional states: attentive and inattentive, 858 and (D) minute by minute. Again, asterisks show the significance values of comparisons examining 859 860 differences as a function of age. (E) Percentage of time infants spent in attentive vs. inattentive states, 861 during the whole interaction. (F) Histogram showing the distribution of the proportion of all the looks that lasted less than or equal to 5 seconds (right) and more than 5 seconds (left) at 5 months (yellow) and 10 862 months (blue). Continuous black line indicates the mode of each distribution. Significance is indicated with 863 864 asterisks where * = p < 0.05, ** = p < 0.01, and *** = p < 0.001. Error bars represent SEMs.

865

Figure 3. ACF and survival probability analyses of the looking behaviour. (**A**) Autocorrelation of the time series looking behaviour at 5 months (in yellow) and 10 months (in blue). (**B**) Survival analysis. Survival probability function for looking behaviour toward object toys. The survival function is the probability that a look will survive a given time. Yellow line shows data from 5-months-old infants with confidence bounds (dotted yellow line) and blue line shows data form 10-months-old infants with confidence bounds (dotted blue line). (**C**) Derivative of the Survival Probability at 5 months (yellow) and 10 months (blue).

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Figure 4. Relationship between infant autonomic arousal and attention. Autocorrelation results for infant autonomic arousal at 5 months (A) and 10 months (C). Significant clusters are indicated by red dots. Cross correlation between infant autonomic arousal and attention at 5 months (B) and at 10 months (D). Infant autonomic arousal forward-predicting infant attention on the negative lags, infant attention forwardpredicting infant autonomic arousal on the positive lags. Black lines show the cross-correlation values, shaded grey areas indicate the SEM. Shaded yellow areas show confidence intervals from the permuted data. Significant time lags identified by the cluster-based permutation analyses are shown by a thick redline.

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Figure 5. Calculation of theta power changes around an attention episode. Results of the linear mixed 883 effects models conducted to examine whether individual looks accompanied by higher theta power are 884 longer lasting. For each look, we calculated the association between the total duration of the look and 885 relative theta power during five time-windows (-2000msec to -1000msec and -1000msec to 0 prior to the 886 look, and 0 to 1000msec, 1000 to 2000msec and 2000 to 3000msec before the look), using a series of 887 separate linear mixed effects models. (A) Shows results at 5 months where the y-axis is the t value, and (B) 888 shows the results at 10 months. Asterisks (*) indicate p values < .05. Central channels include: 'FC1', 'FC2', 889 'C3', 'Cz', 'C4', 'CP1' and 'CP2'. 890

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Figure 6. Calculation of theta power changes within an attention episode. Bar plots for the average
relative theta power throughout a look, at both time points (5 and 10 months) and at different brain networks
(central, occipital, and frontal). Asterisks (*) indicate p<0.05. Error bars represent SEMs.

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Figure 7. Relationship between infant relative theta activity, infant attention, and infant autonomic 896 897 activity. Autocorrelation for infant theta activity at 5 months (A) and at 10 months (D). Significant clusters 898 are indicated by red dots. Cross-correlation results between infant theta activity and infant attention at 5 899 months (B) and at 10 months (E). Infant theta activity forward-predicting infant attention on the negative lags, infant attention forward-predicting infant theta activity on the positive lags. Cross-correlation results 900 901 between infant autonomic activity (indexed by heart rate activity) and relative theta power at 5 months (C) and at 10 months (F). Infant autonomic activity forward-predicting infant theta activity on the negative 902 903 lags, infant theta activity forward-predicting infant autonomic activity on the positive lags. Black lines show the Spearman correlation at each time lag, shaded grey areas indicate the SEM. Shaded yellow areas show 904 905 confidence intervals from the permuted data. Significant time lags identified by the cluster-based permutation analyses are shown by a thick red line. 906



Create new *attention* variable by calculating the duration (msec) of each attentional episode relative to the one-second epochs

Parse heart rate data (analysis 2) & relative theta (analysis 5) into onesecond epochs

Cross-correlate the signals with one-second attention variable



1sec epoch of interest

FFT of each 1s epochs of interest (EOI), for each attentional episode (AE), for each participant



- Extract relative theta (3-6Hz) in central channels
- Separate LMEM (relative power & attentional episode duration) for each of the 5 time windows (EOIs)
- Extract relative theta (3-6Hz) and alpha (6-9Hz) in central, frontal and occipital channels
- Compare the power across the three different time segments









Time window of theta brain activity relative to look onset



