

The effectiveness of psychoanalytic/psychodynamic psychotherapy for reducing suicide attempts and self-harm: Systematic review and meta-analysis

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Abstract

Background

Preventing suicide and self-harm is a global health priority. Although there is a growing evidence-base for the effectiveness of psychoanalytic and psychodynamic psychotherapies for a range of disorders, to date there has been no systematic review of its effectiveness in reducing suicidal and self-harming behaviours.

Aims

To systematically review randomised controlled trials (RCTs) of psychoanalytic and psychodynamic psychotherapies for suicidal attempts and self-harm.

Method

We searched PubMed, PsycINFO, Psycharticles, CINAHL, EMBASE, and the Cochrane Central Register of Controlled Trials for RCTs of psychoanalytic and psychodynamic psychotherapies for reducing suicide attempts and self-harm.

Results

Twelve trials (17 articles) were included in the meta-analyses. Psychoanalytic and psychodynamic therapies were effective in reducing the number of patients attempting suicide (pooled odds ratio = 0.469, 95% CI 0.274 to 0.804). We found some evidence for significantly reduced repetition of self-harm at 6-month follow-up, though not at 12 months.

Significant treatment effects were also found for improvements in psychosocial functioning and reduction in number of hospital admissions.

Conclusions

Psychoanalytic and psychodynamic psychotherapies are indicated to be effective in reducing suicidal behaviour and to have short-term effectiveness in reducing self-harm. They can also be beneficial in improving psychosocial well-being. However, the small number of trials and moderate quality of the evidence means further high quality trials are needed to confirm our findings and to identify which specific components of the psychotherapies are effective.

Declaration of interests

None.

Review registration:

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Introduction

Preventing suicide is a global health priority; it is a significant cause of death in developed and developing countries amongst key population groups, including young people, men in middle age, and people in later life (1). Self-harm and self-injury have also been identified as increasing in many countries over recent years (2). There are more than 100,000 presentations per year at hospitals in England, UK, although this underestimates the numbers: most episodes of self-harm do not reach clinical services and remain in the community (3, 4). There is a strongly evidenced connection between an episode of self-harm or self-injury, repetition and suicide completion; an episode of self-harm increases the chance of suicide completion up to 100 fold (5, 6), thus it is a strong predictor of future suicide attempts (7, 8).

The North American literature distinguishes between suicide attempts and non-suicidal self-injury (NSSI) on the basis of intention to die (9), whilst in the United Kingdom, and Europe the term 'self-harm' is used to include any non-fatal act of harm against the self, irrespective of motivation or method, and so the term 'suicide attempt' is subsumed within the term 'self-harm' in this literature. (3, 10, 11, 12). In this review we use the UK terminology, self-harm, as defined above to include suicide attempts, whilst reporting the terminology used by trial authors, including NSSI and the older term, 'parasuicide'.

Psychological therapies are increasingly important for providing interventions for people at risk of suicide or repeated self-harm, but good quality evidence, especially from Randomised Controlled Trials (RCTs) for the effectiveness of interventions remains low, despite recently increased numbers of trials in this area (3, 10, 11). It is important to know how specific interventions reduce risks and episodes, and which are effective with which groups of people. This review seeks to evaluate the effectiveness of one intervention, psychoanalytic (or psychodynamic) psychotherapy. Psychoanalytic psychotherapy (PP) has a misleading reputation for not being 'evidence based', a reflection of a historic view that it is not amenable to scientific evaluation. A number of recent systematic reviews and meta-analyses attest to the existence of trials evaluating its effectiveness with regard to a range of conditions, including depression, anxiety, somatoform disorders, borderline personality disorder, and for different approaches, including longer and short term therapy, for adults and children and adolescents (13, 14, 15, 16, 17, 18, 19). Conclusions from these reviews are that effect sizes are largely similar to those of other therapies, such as cognitive behavioural therapy, whilst an emerging feature of trials of psychoanalytic psychotherapy is that benefits are sustained and can demonstrate superiority at longer-term follow up (20, 21). This is the first systematic review and meta-analysis of randomised controlled trials evaluating this form of intervention for suicidal behaviour, self-harm and self-injury. The review seeks to assess whether

psychoanalytic psychotherapy is effective in preventing suicide through reducing self-harm (including suicide attempts), reducing symptoms and characteristics linked with known risk factors for suicide/self-harm, and to assess possible adverse effects associated with these interventions¹. To be consistent with recommendations from prior reviews (12) we examine effects for suicide attempts, NSSI and self-harm separately and we explore potential mediators of treatment effects, including type and duration of treatment.

Method

Studies were identified through searches in 6 bibliographical databases (from January 1970 to March 2017): PubMed, PsycINFO, Psycharticles, CINAHL, EMBASE, and the Cochrane Central Register of Controlled Trials. Search terms used were: “psychoanalytic psychotherapy”, “mentalisation based therapy”, “transference-focused psychotherapy”, “cognitive analytic therapy”, “psychodynamic psychotherapy” with suicid*, “self-harm”, “self-mutilation” and “self-injury”. The reference lists of all relevant papers known to the investigators were checked, as were the reference lists of major reviews which were known to the investigators or found by the electronic searches: (10, 11, 13, 14, 17, 18, 22, 23, 24, 25, 26; 27)

Insert Figure 1: Prisma Flow Diagram

¹ The protocol is registered with PROSPERO, CRD42018108483

Studies included had to be RCTs (including cluster randomised and cross-over trials) comparing psychoanalytic psychotherapy of suicidal behaviour, self-harm or self-injury with treatment as usual, routine psychiatric care, enhanced usual care, placebo, or any other comparison, including with a different psychological therapy. Samples included children, adolescents and adults in any setting. The primary outcome measure in this review was the occurrence of repeated self-harm, which, applying the NICE definition means we will assess both suicide attempts and self-injuries (3), including follow up at intervals up to 18 months post-treatment. Secondary outcomes were depression, anxiety, psychosocial functioning and hospital admissions. Where a trial measured the same outcome (for example, depression) in two or more ways, we used the most common measure across trials in any meta-analysis, but we also report scores from the other measures in the text of the review.

We defined the intervention as therapy that was psychoanalytic or psychodynamic in nature, of any duration, aimed at reducing or preventing repetition of suicide attempts, self-harm and self-injury, and reducing suicidal intent, ideation and thoughts.

Data from included trials were extracted by two authors (AG and SB) independently of each other, using a standardised extraction form (adapted from the Cochrane Data Extraction & Assessment Template, 2011). Where the same study had more than one article written about the outcomes, we treated all articles as one study and presented the results only once. We used data from all identified and selected trials to draw funnel plots (size of study versus effect size) (28), to attempt to detect the possibility of publication bias

Statistical analysis

Standardised mean differences were used for continuous data (SMD) and odds ratios (OR) for binary data and accompanying 95% confidence intervals using the random effects

DerSimonian and Laird method. For sub-group comparisons we used fixed effects models. Before groups are compared we first synthesized the results within the groups before comparing at the group level. All analyses were conducted using Stata 14.2. Heterogeneity between studies was measured using the I^2 statistic. We used the following values to denote relative importance of heterogeneity: 0% to 40% (unimportant), 30% to 60% (moderate), 50% to 90% (substantial), and 75% to 100% (considerable), as is the guidance in the Cochrane Handbook (29). Where the value fell into one of the overlapping region (e.g. between 30% to 40%) we used the confidence interval for I^2 or P-value from the Chi^2 test to determine the category of importance. We used the Grading of Recommendations and Evaluation (GRADE) approach to summarise the quality of evidence.

Results

We identified 3290 citations using the search strategy described above. Additionally, 14 papers were identified through searching references. After removing duplicates, the overall number of citations remaining was 2919. Of these, 2894 were excluded after screening and a further 9 were excluded after reading full texts. The 17 papers included reported from 12 studies. All studies included were described as randomised controlled trials. All but one study used a simple randomisation procedure, based on individual allocation to the intervention and control groups. One trial (30) randomised clinicians in blocks rather than participants. The main characteristics of the studies are summarised in online Table 1 (included studies) and Table 2 (excluded studies). Eight excluded studies are listed in online Table 2. Three were excluded because the research design was not RCT (31; 32, 33); two were because the intervention was judged not to be psychodynamic (34; 35); three were excluded because outcomes for self-harm were not assessed (36, 37, 38)

Description of Studies

12 studies were included (30, 39-54). Characteristics of the trials are shown in the online studies table (Table 1). The total number of randomised participants was 999. Some studies excluded some drop-outs in the data analysis (45, 46, 48, 50, 52, 53) and thus the total participants analysed was 939. Samples ranged from 30 (24 after excluding drop outs in the analysis) to 134, with means of 76.8 randomised, and 72.2 analysed.

Studies were conducted in outpatient and community settings (9 studies), a combination of A&E and outpatient (30), A&E and patients' homes (51), and an inpatient setting (40). Studies were conducted in several countries: the UK (30, 40, 44, 55, 51, 54), USA (46, 50), Europe (39, 47, 48, 53) and Australia (45). Nine studies were of adult samples and three (30, 45, 54) were of adolescents under 18 years. All studies excepting one (39) had a majority of female participants, and one study (47) recruited only females. Overall, female participants outnumbered males by a ratio of 4:1, which happens to reflect self-harm gender disparities found in the prevalence literature.

A range of controls was used in these studies; no waiting list or cross over designs were used. TAU, of various kinds, was used in 6 trials (30, 40, 47, 51, 53, 54) and enhanced usual care in 4 studies (39, 44, 45, 50). Two trials compared PP with another psychological therapy: with schema-focused therapy (48), and with dialectical behaviour therapy and supportive therapy (46).

In all included studies, the intervention employed core psychoanalytic methods to increase awareness and self-reflection; to manage, regulate or contain emotions; and to effect change through the therapeutic relationship. However, there were differences of approach, intensity and duration of the intervention and differences between types of study population (age,

number of previous episodes of self-harm). We identified subgroups for the purposes of making comparisons; for population (adult or adolescent); for acute or chronic self-harm; duration of therapy (short term <40 sessions or longer term > 40 sessions), and for comparators (TAU, enhanced TAU or another intervention). Additionally, we assessed the experience of therapists and their training in and adherence to the intervention.

Primary Outcomes

Suicide attempts

There was an overall significant treatment effect for PP compared to treatment as usual on the number of patients who attempted suicide in three studies at 12-month follow-up (3 studies; n = 276, pooled OR = 0.469, 95% CI 0.274 to 0.804, $I^2 = 0.0$) (Figure 2). There was no significant treatment effect for PP on the number of suicide attempts (episodes) at 12-month follow up (2 studies; n = 238, SMD = -0.235, 95% CI: = -0.502 to 0.033, $I^2=8.6$) (Figure 3).

Self-harm

There was evidence of a significant treatment effect for PP on the number of patients who repeated self-harm (i.e. self-harm and suicide attempts combined) by the 6-month follow up (2 studies; n=125, OR=0.27, 95% CI=0.109 to 0.668, p=0.005, $I^2=82.7\%$) (Figure 4). High heterogeneity should be noted. However, at 12-month follow up there was no evidence of a significant treatment effect for PP on the number of patients who repeated self-harm (3 studies; n=278, OR = 0.581, 95%, CI = 0.236 to 1.426, $I^2 = 63.6\%$). Similarly, there was no evidence of significant treatment effect for PP on self-harm episodes in this period (rather than the number of patients who repeated self-harm) at 12-month follow up on the basis on data from three studies (n=273; SMD = -0.149, 95% CI = 0.388 to 0.089, $I^2 = 0.0\%$) (Figure 5).

Secondary outcomes

Depression

Data on depression scores using the Beck Depression Inventory (BDI) at six months' follow-up did not suggest a significant treatment effect for PP (two studies; $n=172$, $SMD = -0.273$ 95% CI = -0.615 to 0.069 , $I^2 = 0.0\%$). There was no overall evidence to suggest a significant treatment effect for PP at 12-month follow-up (3 studies; $n = 172$, $SMD = -0.357$, 95% CI = -1.048 to 0.335 , $I^2 = 85.2\%$).

Anxiety

There was no overall evidence of a significant treatment for PP at 12-month follow-up in anxiety state (temporary state of anxiety) assessed by the Spielberger State-trait Anxiety Inventory (2 studies; $n = 142$, $SMD = -0.350$, 95% CI = -1.458 to 0.757 , $I^2 = 87.8\%$). It ought to be noted that significant heterogeneity was found. As for anxiety trait (long-term anxiety), there was no evidence to suggest a treatment effect for PP at 12 months follow-up (2 studies; $n=142$, $SMD = 0.018$, 95% CI = -0.311 to 0.347 , efficacy at 12 months, $I^2 = 0.0\%$).

Psychosocial Functioning

Psychosocial functioning was assessed by the Social Adjustment Scale (SAS). There was evidence of an overall moderate treatment effect for PP at the 18-month follow-up assessment (3 studies; $n = 207$, $SMD = -0.620$, 95% CI = -1.210 to -0.030 , $I^2 = 23.6\%$).

However, PP was not found to have a significant effect on psychosocial functioning when that was assessed by the Global Assessment of Functioning Scale (GAS) in two trials at 12-months follow up (2 studies; $n = 139$, $SMD = 0.269$, 95% CI = 0.068 to 0.606 , $I^2 = 0.0\%$). Overall, PP did not show an effect on improved psychosocial functioning as assessed by GAS at 18-month follow-up (two studies; $n = 169$, $SMD = 0.401$, 95% CI = -0.034 to 0.837 , $I^2 = 34.5\%$).

Admissions

There was evidence of a significant treatment effect for PP on the reduction of number of admissions at the 12-month follow-up assessment (two studies, $n = 238$, $SMD = -0.505$, 95% $CI = -0.763$ to -0.246 , $I^2 = 0.0\%$). However, there was no evidence of a treatment effect for PP in reducing the number of admissions at six-month follow-up in two trials ($n = 253$, $SMD = -0.291$, 95% $CI = -0.592$ to 0.011 , $I^2 = 23.6\%$). Pooled analysis of three trials at 12-month follow-up assessment suggested no overall evidence of a significant treatment effect for PP in reducing the length of admissions in days ($n = 273$, $SMD = -0.268$, 95% $CI = -0.563$ to 0.028 , $I^2 = 27.9\%$).

Differences between subgroups

We found no evidence of significant differences between studies of different populations (adult or adolescent), self-harm histories (acute or chronic) or for long and short-term interventions ($>$ or $<$ treatment for one year). In all comparisons, heterogeneity was low ($I^2 < 30\%$). For comparisons of studies using different comparators (TAU and enhanced TAU) heterogeneity was also low ($I^2 < 30\%$).

Quality of the Studies

Bias was detected in the domain of performance and detection bias, with 5 trials being at high risk for blinding of participants, though this can be problematic in trials of this nature. Overall we judged the majority of the 12 studies to be at low risk of bias. There was inadequate information to make a satisfactory risk of bias judgement for the majority of trials (9/12) for

selected reporting bias (for detailed ratings see the online Table 3). Using funnel plots we checked for publication bias in three groups of outcomes: those reported as means, those reporting counts and those with binary outcomes. In all cases the plots revealed significant asymmetry suggesting publication bias. This was confirmed by the Egger test (studies reporting means: bias = -4.86 t= -3.04 p= 0.008; studies reporting counts: bias 2.74, t = 2.83, p = 0.016; and studies reporting binary outcomes: bias = -1.78, t= -2.10, p = 0.048) (Figure 6). Applying the GRADE system we classified the quality of evidence overall as moderate.

All studies used a manual/protocol. Delivering therapists ranged from expert (2 studies), moderate training and experience (8 studies) to low experience and qualifications (2 studies). In 6 studies training to use the intervention was high; 3 studies were moderate and 3 studies low; supervision and fidelity measures were high in 7 studies, moderate in 4 and low in 1.

Discussion

This systematic review seeks to provide an estimation of the effectiveness of psychoanalytic psychotherapies compared with treatment as usual and with other interventions for reducing suicide attempts and self-harm. This included the evaluation of whether such interventions had a positive effect on proximal risk factors for suicide attempts and self-harm such as depression and anxiety, for reduction of psychosocial problems, and for reduction of length of hospital admissions associated with suicide and self-harm. We also evaluated the quality of the evidence base and the presence of publication bias.

The finding in relation to the primary outcomes was that there was some evidence for a significant treatment effect for the number of patients attempting suicide (pooled OR = 0.468 with 276 participants) at 12-month follow-up. There was also some evidence for a small

significant effect for the reduction of self-harm (defined here to include suicide attempts and self-injuries) at 6-month follow up (OR=0.27 with 149 participants), and it is noted there is high heterogeneity for this result. There was no evidence for a significant treatment effect for self-harm or suicidal episodes, nor the number of patients repeating self-harm at 12-month follow up. Though other studies of PP show that differences are sustained at longer-term follow up we were not able to pool findings from studies for the effects at longer term, i.e. 18 months and longer. The clinical implication is that PP has benefit over comparators for self-harm 6 months after treatment ends, but not at 12 months; for suicide attempts the evidence is in favour of the intervention at 12 months.

We did not find overall a significant treatment effect for improvement of depression or anxiety symptoms in studies of PP at 6 months and 12 months follow-up. The findings were similar in relation to both anxiety state (temporary anxiety) and anxiety trait (long-term anxiety) with no overall significant treatment effect for psychodynamic therapies at either 12 months or 18 months.

We looked for improvements in psychosocial functioning in patients who received psychoanalytic psychotherapy. Studies use different outcome measures so data could only be pooled where the same measurement instrument had been used. Studies that used the Social Adjustment Scale found evidence of an overall moderate treatment effect at 18 month follow up, based on three studies with SMD -0.620 and 207 participants. However, those studies which used GAS as the measurement tool for psychosocial functioning found no overall significant treatment effect.

Finally, we considered whether PP reduced the number of hospital admissions, or the length of admissions. There was no effect in reducing the number of admissions at six months, though

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there was a significant effect at 12-month follow-up (2 studies; SMD = -0.505 with 238 participants). There was no overall evidence of reduction in the number of days spent in hospital at twelve-month follow-up.

We explored how differences of population, histories of self-harm, types of comparator, types of psychodynamic therapy, affected the results. We found no significant heterogeneity between studies for different populations, in terms of age (adult v adolescent), self-harm history (acute v chronic) and duration of therapy (> or < one year) Despite different nomenclatures, all the interventions applied psychoanalytic principles to improve awareness, emotion regulation and relationships and to effect change through the therapeutic relationship.

Quality of the Evidence

Risk of bias was unclear for several domains in most of the studies, particularly in relation to reporting bias (selective reporting) where 9 of the included 12 studies were rated as 'unclear'. Risk of bias was considered low for the significant majority of studies in relation to selection bias, performance and detection bias (with the exception of blinding of participants, which has inherent difficulties in researching psychotherapeutic interventions), and attrition bias. Applying the GRADE system we classified the quality of evidence as moderate. The possibility of significant publication bias was found, though the funnel plot asymmetry could be due to the different populations sampled in the studies. Our analyses were based on a limited number of trials overall and we were unable to conduct other sensitivity analyses because of lack of power of potential comparisons.

Limitations

The review is based on a limited number of studies. The interventions that were investigated in the trials, although selected for inclusion because they met the descriptor for psychoanalytic

psychotherapy, were diverse in relation to intensity, content and duration of treatment. We addressed this through comparing outcomes for subgroups for age, extent of self-harm and duration of therapy and found no significant heterogeneity. However, it is difficult to identify which specific components had a therapeutic effect. This is further complicated by inconsistency of taxonomies of psychoanalytic psychotherapies, reflecting an apparent predisposition in the field for the constant invention of new labels for similar interventions; despite this, all interventions adhered to principles for psychoanalytic psychotherapy. Likewise, divergence in the way outcomes were measured and recorded, and variations in inclusion and exclusion criteria present challenges for the systematic reviewer. This includes the different definitions of self-harm used in various geographical areas. We had hoped to include suicidal ideation in the review but this is not reported in most studies, and in the few studies where it was reported there was no standardisation of measurement, so this was not possible. Finally, studies did not address patients' experiences of therapy.

Implications for practice

Our systematic review of the effectiveness of PP in reducing suicidal and self-harming behaviour provides some preliminary evidence that PP can reduce the number of patients attempting suicide. There is some evidence that suicide attempts are reduced up to 12 months follow up, and self-harm is reduced in the short-term (6 months follow up), but we found no evidence that PP reduces self-harm more than the comparators at 12 months. An implication is that PP can be used to aim to reduce self-harm more effectively up to 6 months after the end of treatment; the finding shows that comparators 'catch up' by 12 months. Further studies with longer-term follow up (18 and 24 months) are needed to understand this better. Where we were able to pool findings at 18 months follow up, for psychosocial functioning, we did

find some evidence of an effect in favour of the intervention, which is consistent with other studies for PP (14, 16).

There is some evidence of effectiveness in some risk factors for suicidal and self-harming behaviour. PP has a positive impact on psychosocial functioning, and thus on reducing problems, and reduces length of inpatient admissions. We note that meta-analyses of psychological therapies for self-harm and suicide reach broadly comparable conclusions, with some evidence for reducing repeated self-harm/suicide attempts and some reduction of associated risk factors together with findings where no evidence is found in favour of the interventions (11). Hence, this review suggests that service providers could consider the use of PP as an intervention that could be offered to individuals at risk of, or with a history of, suicidal or self-harming behaviour. No cost data were available with the trials included in the review, save one (42), but these PP interventions include structured, manualised approaches; in their Cochrane review of short-term psychoanalytic therapies, Abbass et al 2014 (13) note that where psychoanalytic therapies are manualised and short term these can be relatively inexpensive and might be a prudent first-line intervention before more extensive interventions are considered. It must be noted, however, that longer term PP is beginning to demonstrate effectiveness for a number of conditions (16), and in this review (40, 44, 47), and thus its value should not be underestimated.

Future research

Our systematic review and meta-analysis suggest that PP is potentially effective in the treatment of suicidal and self-harming behaviours, and in reduction of some of the risk factors for these behaviours. However, given the relatively small number of RCTs meeting the criteria for inclusion in the analysis, and the presence of bias in these studies, including potential publication bias, there is a need for further high-quality research. Future studies should also

focus on specific diagnostic and problem categories to provide practitioners with evidence with which to consider these treatments for specific populations. Studies should also examine specific treatment factors and their contributions to outcomes across therapy models. Studies should ensure that follow-up is conducted over the longer-term, up to at least 24 months after the end of treatment.

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